

On the Error Bound in the Normal Approximation for  
Jack Measures  
*(Joint work with Le Van Thanh)*

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# Partitions of Positive Integers

- A partition of a positive integer  $n$  is a finite non-increasing sequence of positive integers  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_l > 0$  such that

$$\sum_{i=1}^l \lambda_i = n. \text{ Write } \lambda = (\lambda_1, \lambda_2, \dots, \lambda_l).$$

- The  $\lambda_i$  are called the parts of the partition  $\lambda$  and the number  $l$  of parts called the length of  $\lambda$ .
- We write  $\lambda \vdash n$  to denote " $\lambda$  is a partition of  $n$ ".
- Denote that set of all partitions of  $n$  by  $\mathcal{P}_n$  and the set of all partitions by  $\mathcal{P}$ , that is,  $\mathcal{P} = \bigcup_{n=0}^{\infty} \mathcal{P}_n$ . By convention, the empty sequence forms the only partition of zero.

# Partitions of Positive Integers

- Let  $p(n)$  be the partition function, that is, the number of partitions of  $n$ .
- An important and fundamental question is to evaluate  $p(n)$ .
- Euler started the analytic theory of partitions by providing an explicit formula for the generating function of  $p(n)$ :

$$\mathcal{F}(q) := \sum_{n=0}^{\infty} p(n)q^n = \prod_{k=1}^{\infty} \frac{1}{1 - q^k}.$$

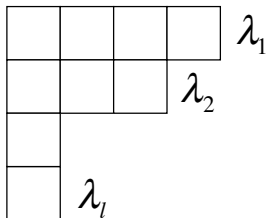
- In a celebrated series of memoirs published in 1917 and 1918, Hardy and Ramanujan established:

$$p(n) = \frac{1}{4n\sqrt{3}} e^{\pi\sqrt{\frac{2}{3}n}} \left(1 + O\left(\frac{1}{\sqrt{n}}\right)\right).$$

# Young Diagram

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- To each partition  $\lambda = (\lambda_1, \lambda_2, \dots, \lambda_l)$  is associated its Young diagram (shape).



# Young Tableau

- A standard Young tableau  $T$  with the shape  $\lambda \vdash n$  is a one-to-one assignment of the numbers  $1, 2, \dots, n$  to the squares of  $\lambda$  in such a way that the numbers increase along the rows and down the columns. See, for example,  $n = 9$ .

1	3	4	7	$\lambda_1$
2	5	6	$\lambda_2$	
8				
9	$\lambda_1$			

- Let  $d_\lambda$  denote the total number of standard Young tableaux associated with a given shape  $\lambda$ .

# Plancherel Measure

- The set of irreducible representations of the symmetric group  $S_n$  of permutations of  $1, 2, \dots, n$  can be parameterized by  $\lambda \in \mathcal{P}_n$ .
- The degree (dimension) of the irreducible representation indexed by  $\lambda$  is equal to  $d_\lambda$ .
- The Burnside identity is:

$$\sum_{\lambda \vdash n} d_\lambda^2 = n! \quad (\text{that is, } \sum_{\lambda \vdash n} \frac{d_\lambda^2}{n!} = 1).$$

- The Plancherel measure is a probability measure on  $\lambda \vdash n$  (also on the irreducible representations of  $S_n$ , parameterized by  $\lambda$ ) given by:

$$P(\{\lambda\}) = \frac{d_\lambda^2}{n!}.$$

# Plancherel Measure

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- The first row of a random partition distributed according to the Plancherel measure has the same distribution as the longest increasing subsequence of a random permutation distributed according to the uniform measure.
- Let  $l(\pi)$  be the length of the longest increasing subsequence of the random permutation  $\pi$ . It is known that  $(l(\pi) - 2\sqrt{n})/n^{1/6}$  converges to the Tracy-Widom distribution. (Baik, Deift and Johansson (1999), *J. Amer. Math. Soc.*)



## Character Ratio

- The character of a group representation is a function on the group that associates to each group element the trace of the corresponding matrix. It is called irreducible if it is the character of an irreducible representation.
- Let  $\chi^\lambda(12)$  be the irreducible character parametrized by  $\lambda$  evaluated on the transposition  $(12)$ .
- The quantity  $\frac{\chi^\lambda(12)}{d_\lambda}$  is called a character ratio.
- The eigenvalues for the random walk on the symmetric group generated by transpositions are the character ratios  $\chi^\lambda(12)/d_\lambda$ , each occurring with multiplicity  $d_\lambda^2$ . [Diaconis and Shahshahani \(1981\), \*Z. Wahr. Verw. Gebiete\*](#).
- Character ratios also play an essential role in work on the moduli spaces of curves (see [Eskin and Okounkov \(2001\), \*Invent. Math.\*](#) and [Okounkov and Pandharipande \(2005\), \*Proc. Sympos. Pure Math., 80, Part 1, Amer. Math. Soc.\*](#)).

# Normal Approximation for Character Ratios

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Let

$$W_n = \sqrt{\binom{n}{2}} \frac{\chi^\lambda(12)}{d_\lambda}$$

and let  $\Phi$  be the  $\mathcal{N}(0, 1)$  distribution function. Assume  $n \geq 2$  and let  $x \in \mathbb{R}$ .

- Kerov (1993), *Compt. Rend. Acad. Sci. Paris*.

$$W_n \xrightarrow{\mathcal{L}} \mathcal{N}(0, 1) \text{ as } n \rightarrow \infty.$$

- Fulman (2005), *Trans. AMS* (using Stein's method)

$$|P(W_n \leq x) - \Phi(x)| \leq 40.1n^{-1/4}.$$

# Normal Approximation for Character Ratios

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- Fulman (2006) *Trans. AMS* (using martingales)

$$|P(W_n \leq x) - \Phi(x)| \leq Cn^{-s} \text{ for any } s < 1/2.$$

- Shao and Su (2006), *Proc. AMS* (using Stein's method)

$$|P(W_n \leq x) - \Phi(x)| \leq Cn^{-1/2}.$$

# Jack Measures

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The  $\text{Jack}_\alpha$  measure,  $\alpha > 0$ , is a probability measure on  $\lambda \vdash n$  given by:

$$\text{Jack}_\alpha(\lambda) = \frac{\alpha^n n!}{\prod_{x \in \lambda} (\alpha a(x) + l(x) + 1)(\alpha a(x) + l(x) + \alpha)},$$

where in the product over all boxes  $x$  in the partition  $\lambda$ ,

(i)  $a(x)$  denotes the number of boxes in the same row of  $x$  and to the right of  $x$  (the "arm" of  $x$ ),

(ii)  $l(x)$  denotes the number of boxes in the same column of  $x$  and below  $x$  (the "leg" of  $x$ ).

## Jack Measures

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For example, take  $n = 5$  and  $\lambda$  as shown below.

$$\lambda = \begin{array}{|c|c|c|} \hline & & \\ \hline & & \\ \hline & & \\ \hline & & \\ \hline \end{array}$$

$$\begin{aligned} \text{Jack}_\alpha(\lambda) &= \frac{\alpha^n n!}{\prod_{x \in \lambda} (\alpha a(x) + l(x) + 1)(\alpha a(x) + l(x) + \alpha)} \\ &= \frac{60\alpha^2}{(2\alpha + 2)(3\alpha + 1)(\alpha + 2)(2\alpha + 1)(\alpha + 1)}. \end{aligned}$$

# Jack Measures

- The  $\text{Jack}_\alpha$  measure with  $\alpha = 2/\beta$  is a discrete analog of Dyson's  $\beta$  ensembles in random matrix theory, which are tractable for  $\beta = 1, 2, 4$ .
- The joint probability density for the eigenvalues  $x_1 \geq x_2 \geq \cdots \geq x_n$  of the Gaussian orthogonal ensemble (GOE), Gaussian unitary ensemble (GUE) and Gaussian symplectic ensemble (GSE) is given by

$$\frac{1}{Z_\beta} \exp\left(-\frac{x_1^2 + \cdots + x_n^2}{2}\right) \prod_{1 \leq i < j \leq n} (x_i - x_j)^\beta$$

for  $\beta = 1, 2, 4$  respectively.

- The  $\text{Jack}_\alpha$  measure with  $\alpha (= 2/\beta) = 2, 1, 1/2$  has group theoretical interpretation.

# Jack Measures

- In the case  $\alpha = 1$ ,

$$\begin{aligned} \text{Jack}_\alpha(\lambda) &= \frac{\alpha^n n!}{\prod_{x \in \lambda} (\alpha a(x) + l(x) + 1)(\alpha a(x) + l(x) + \alpha)} \\ &= \frac{n!}{\prod_{x \in \lambda} h^2(x)}, \end{aligned}$$

where  $h(x) = a(x) + l(x) + 1$  is the hook length of the box  $x$ .

- The hook-length formula states that

$$d_\lambda = \frac{n!}{\prod_{x \in \lambda} h(x)}.$$

- Hence the Plancherel measure can be expressed as

$$P(\{\lambda\}) = \frac{d_\lambda^2}{n!} = \frac{n!}{\prod_{x \in \lambda} h^2(x)},$$

which agrees with the  $\text{Jack}_\alpha$  measure for  $\alpha = 1$ .

# Normal Approximation for Jack Measures

Let

$$W_{n,\alpha} = W_{n,\alpha}(\lambda) = \frac{\sum_i \left( \alpha \binom{\lambda_i}{2} - \binom{\lambda'_i}{2} \right)}{\sqrt{\alpha \binom{n}{2}}},$$

where the partition  $\lambda \vdash n$  is chosen according to the  $\text{Jack}_\alpha$  measure,  $\lambda_i$  is the length of the  $i$ th row of  $\lambda$  and  $\lambda'_i$  the length of the  $i$ th column of  $\lambda$ .

If  $\alpha = 1$ ,

$$W_{n,\alpha} = \sqrt{\binom{n}{2}} \frac{\chi^\lambda(12)}{d_\lambda}$$

by the Frobenius formula.



# Normal Approximation for Jack Measures

Assume  $n \geq 2$  and let  $x \in \mathbb{R}$ .

- Fulman (2004), *J. Comb. Theory Ser. A*

For  $\alpha \geq 1$ ,

$$|P(W_{n,\alpha} \leq x) - \Phi(x)| \leq \frac{C_\alpha}{n^{1/4}}.$$

He conjectured that for  $\alpha \geq 1$ , the optimal bound is a universal constant multiplied by  $\max\left\{\frac{1}{\sqrt{n}}, \frac{\sqrt{\alpha}}{n}\right\}$ .

- Fulman (2006) *Trans. AMS* (using martingales)

For  $\alpha \geq 1$ ,

$$|P(W_{n,\alpha} \leq x) - \Phi(x)| \leq \frac{C_\alpha}{n^{1/2-\epsilon}} \quad \text{for any } \epsilon > 0.$$

# Normal Approximation for Jack Measures

- Fulman (2006), *Ann. Comb.* (using Stein's method)

For  $\alpha \geq 1$ ,

$$|P(W_{n,\alpha} \leq x) - \Phi(x)| \leq \frac{C_\alpha}{n^{1/2}}.$$

- Fulman and Goldstein (2011), *Comb. Probab. Comput.* (using Stein's method and zero-bias coupling)

For  $\alpha > 0$ ,

$$\|F - \Phi\|_1 \leq \sqrt{\frac{2}{n}} \left( 2 + \sqrt{2 + \frac{\max(\alpha, 1/\alpha)}{n-1}} \right),$$

where  $F(x) = P(W_{n,\alpha} \leq x)$ .

Chen and Thanh (2014), *Preprint*

## Theorem 1

For  $\alpha > 0$ ,

$$\sup_{x \in \mathbb{R}} |P(W_{n,\alpha} \leq x) - \Phi(x)| \leq 9 \max \left\{ \frac{1}{\sqrt{n}}, \frac{\max\{\sqrt{\alpha}, 1/\sqrt{\alpha}\} \log n}{n} \right\}.$$

### Remarks.

1. For  $\alpha = 1$ , the theorem reduces to one for character ratios under the Plancherel measure with the bound  $\frac{9}{\sqrt{n}}$ , where the constant is explicit.

2. For  $\alpha \geq 1$ , the bound becomes  $9 \max \left\{ \frac{1}{\sqrt{n}}, \frac{\sqrt{\alpha} \log n}{n} \right\}$ , which is close to that conjectured by [Fulman \(2004\)](#).

Chen and Thanh (2014), *Preprint*

## Theorem 2

For  $\alpha > 0$  and  $p \geq 2$ , and for  $x \in \mathbb{R}$ ,

$$|P(W_{n,\alpha} \leq x) - \Phi(x)| \leq \frac{C_p}{1 + |x|^p} \max \left\{ \frac{1}{\sqrt{n}}, \frac{\max\{\sqrt{\alpha}, 1/\sqrt{\alpha}\} \log n}{n} \right\}$$

where  $C_p$  is a constant depending only on  $p$ .

## Remarks.

1. For  $\alpha = 1$ , the theorem reduces to one for character ratios under the Plancherel measure with the bound  $\frac{C_p}{1 + |x|^p} \frac{1}{\sqrt{n}}$ .
2. For  $\alpha \geq 1$ , the bound becomes  $\frac{C_p}{1 + |x|^p} \max \left\{ \frac{1}{\sqrt{n}}, \frac{\sqrt{\alpha} \log n}{n} \right\}$ .

# Zero-bias Coupling

Goldstein and Reinert (1997), *Ann. Appl. Probab.*

- For  $W$  with  $EW = 0$  and  $\text{Var}(W) = B^2$ , there always exists  $W^*$  such that  $EWf(W) = B^2Ef'(W^*)$  for absolutely continuous functions  $f$  for which the expectations exist.
- The distribution of  $W^*$  is called  $W$ -zero-biased.
- $W^*$  must necessarily be absolutely continuous and its density function is given by  $B^{-2}EWI(W > x)$ .
- Not easy to find couplings of  $W$  with  $W^*$  which are effective for normal approximation.
- Effective couplings are known for  $W = \sum_{i=1}^n X_i$ , where the  $X_i$  are independent and for  $W = \sum_{i=1}^n a_{i\pi(i)}$  where  $\pi$  is a random permutation (Goldstein (2005), *J. Appl. Probab.*).

# Rosenthal Inequality for Zero-bias Coupling

Chen and Thanh (2014), *Preprint*

## Proposition 3

Let  $W$  be such that  $EW = 0$  and  $\text{Var}(W) = B^2 > 0$ . Suppose  $W$ -zero-biased  $W^*$  is defined on the same probability space as  $W$ . Then for  $p \geq 2$ ,

$$E|W|^p \leq \kappa_p \{B^p + B^2 E|W^* - W|^{p-2}\}$$

where  $\kappa_p = 2^{(p-2)+(p-4)+\dots}(p-1)(p-3)\dots$

If  $W = \sum_{i=1}^n X_i$ , where the  $X_i$  are independent with zero mean, then

$$E|W|^p \leq \kappa_p B^p + \bar{\kappa}_p \sum_{i=1}^n E|X_i|^p$$

where  $\bar{\kappa}_p = 2 \max\{1, 2^{p-3}\} \kappa_p$ .

# Normal Approximation for Zero-bias Coupling

Chen and Thanh (2014), *Preprint*

## Theorem 4

Let  $EW = 0$  and  $\text{Var}(W) = 1$ . Suppose the zero-biased  $W^*$  is defined on the same probability space as  $W$ . Let  $T = W^* - W$ .

1. Then

$$\sup_{x \in \mathbb{R}} |P(W^* \leq x) - \Phi(x)| \leq \sqrt{ET^2} + \frac{\sqrt{2\pi}}{4} E|T|$$

2. Assume  $E|T|^{2p} \leq 1$  for some  $p \geq 2$ . Then for all  $x \in \mathbb{R}$ ,

$$|P(W^* \leq x) - \Phi(x)| \leq \frac{C_p(\sqrt{ET^2} + \sqrt{ET^4} + E|T|^{p+1} + E|T|^{p+2})}{1 + |x|^p}.$$

# Normal Approximation for Zero-bias Coupling

## Sketch of proof of Theorem 4 part 2

Since  $(-W)^* = -W^*$  and in view of part 1, it suffices to assume  $x \geq 2$ . Using the properties of the solution of the Stein equation,

$$|P(W^* \leq x) - \Phi(x)| \leq \frac{C_p(E|T| + E|T|^{p+1} + E|T|^{p+2})}{1 + x^p} \\ + \frac{C_p(\sqrt{E|W|^{2p}E|T|^2} + \sqrt{E|W|^{2p+2}E|T|^2} + \sqrt{E|W|^{2p}E|T|^4})}{1 + x^p}.$$

Since  $E|T|^{2p} \leq 1$ , by the Rosenthal inequality,

$$E|W|^{p+2} \leq C_p(1 + E|T|^{2p}) \leq C_p.$$

Similarly,

$$E|W|^{2p} \leq C_p.$$



## Zero-bias Coupling for Jack Measures

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Kerov's growth process (Kerov (2000), *Funct. Anal. Appl.*) gives a sequence of partitions  $(\lambda(1), \lambda(2), \dots, \lambda(n))$ , where for each  $j$ ,  $\lambda(j)$  is a partition of  $j$  distributed according to the  $\text{Jack}_\alpha$  measure. Using this process, one can show that

$$W_{n,\alpha}(\lambda) = \frac{\sum_{x \in \lambda} c_\alpha(x)}{\sqrt{\alpha \binom{n}{2}}}$$

where  $c_\alpha(x)$  denotes the " $\alpha$ -content" of  $x$ , which is defined as

$$c_\alpha(x) = \alpha[(\text{column number of } x) - 1] - [(\text{row number of } x) - 1].$$

## Zero-bias Coupling for Jack Measures

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In the diagram below, representing a partition of 7, each box is filled with its  $\alpha$ -content.

Recall that

$$c_\alpha(x) = \alpha[(\text{column number of } x) - 1] - [(\text{row number of } x) - 1].$$

0	$1\alpha$	$2\alpha$	$3\alpha$
-1	$\alpha-1$		
-2			

## Zero-bias Coupling for Jack Measures

Fulman and Goldstein (2011), *Comb. Probab. Comput.*,

The statistic  $W_{n,\alpha}$  and its zero-biased  $W_{n,\alpha}^*$  are coupled as follows:

$$W_{n,\alpha} = V_{n,\alpha} + \eta_{n,\alpha} \quad \text{and} \quad W_{n,\alpha}^* = V_{n,\alpha} + \eta_{n,\alpha}^*$$

where  $\eta_{n,\alpha}^*$  is  $\eta_{n,\alpha}$ -zero-biased,  $V_{n,\alpha}$ ,  $\eta_{n,\alpha}$  and  $\eta_{n,\alpha}^*$  are defined on the same probability space,

$$V_{n,\alpha} = \sum_{x \in \nu} c_\alpha(x) / \sqrt{\alpha \binom{n}{2}} = \sqrt{\frac{n-2}{n}} W_{n-1,\alpha},$$

$$\eta_{n,\alpha} = c_\alpha(\lambda/\nu) / \sqrt{\alpha \binom{n}{2}},$$

$\nu$  is a partition of  $n-1$  chosen from the  $\text{Jack}_\alpha$  measure, and  $c_\alpha(\lambda/\nu)$  denotes the  $\alpha$ -content of the box added to  $\nu$  to obtain  $\lambda$ .

Some moment bounds on  $\eta_{n,\alpha}$  are also obtained.

# Sketch of Proof of Main Theorems

Chen and Thanh (2014), *Preprint*

## Lemma 5

For  $p \geq 1$  and  $\alpha \geq 1$ ,

$$P \left( |\eta_{n,\alpha}| \geq \frac{p\sqrt{2e^3}}{\sqrt{n-1}} \right) \leq \frac{n}{2\pi(p^2e)^{p\sqrt{e^3n/\alpha}}},$$

$$P \left( |\eta_{n,\alpha}^*| \geq \frac{p\sqrt{2e^3}}{\sqrt{n-1}} \right) \leq \frac{\alpha n^2}{2\pi(p^2e)^{p\sqrt{e^3n/\alpha}}}.$$

Recall that  $W_{n,\alpha}^* - W_{n,\alpha} = \eta_{n,\alpha}^* - \eta_{n,\alpha}$ .

# Sketch of Proof of Main Theorems

## Theorem 6

Let  $EW = 0$  and  $\text{Var}(W) = 1$ . Suppose the zero-biased  $W^*$  is defined on the same probability space as  $W$ . Let  $T = W^* - W$  and let  $\epsilon \geq 0$ .

1. Then

$$\sup_{x \in \mathbb{R}} |P(W \leq x) - \Phi(x)| \leq \sqrt{ET^2} + \frac{\sqrt{2\pi}}{4} E|T| + \frac{\epsilon}{\sqrt{2\pi}} + P(|T| > \epsilon).$$

2. Assume  $E|T|^{2p} \leq 1$  for some  $p \geq 2$ . Then for all  $x \in \mathbb{R}$ ,

$$|P(W \leq x) - \Phi(x)| \leq \frac{C_p(\sqrt{ET^2} + \sqrt{ET^4} + E|T|^{p+1} + E|T|^{p+2})}{1 + |x|^p} + \frac{\epsilon + \sqrt{P(|T| > \epsilon)}}{1 + |x|^p}.$$

# Sketch of Proof of Main Theorems

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- Theorem 6 is deduced from Theorem 4.
- Combine Lemma 5 and Theorem 6 to prove Theorem 1 and Theorem 2 for  $\alpha \geq 1$ .
- For  $0 < \alpha < 1$ , note that from the definition of the  $\text{Jack}_\alpha$  measure,  $P_\alpha(\lambda) = P_{1/\alpha}(\lambda^t)$ , where  $\lambda^t$  is the transpose partition of  $\lambda$ .
- Also from its definition,  $W_{n,\alpha}(\lambda) = -W_{n,1/\alpha}(\lambda^t)$ .

# Summary

- The  $\text{Jack}_\alpha$  measure on partitions of a positive integer is a discrete analog of Dyson's  $\beta$  ensembles in random matrix theory.
- For  $\alpha = 1$ , the  $\text{Jack}_\alpha$  measure agrees with the Plancherel measure on the irreducible representations of the symmetric group.
- We obtained both uniform and non-uniform Berry-Esseen bounds for  $W_{n,\alpha} = \frac{\sum_i (\alpha \binom{\lambda_i}{2} - \binom{\lambda'_i}{2})}{\sqrt{\alpha \binom{n}{2}}}$ , where  $\alpha > 0$  and the partition  $\lambda$  is chosen from the  $\text{Jack}_\alpha$  measure. If  $\alpha = 1$ ,  $W_{n,\alpha}$  coincides with  $\sqrt{\binom{n}{2}} \frac{\chi^\lambda(12)}{d_\lambda}$ .
- For  $\alpha \geq 1$ , we came close to solving a conjecture of [Fulman \(2004\)](#).

Thank You