

# MATH 505a Spring 2021 Qual Solution Attempts

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## Problem 1

A permutation  $\pi$  on  $n$  symbols is said to have  $i$  as a fixed point if  $\pi(i) = i$ .

(a)

Find the probability  $p_n$  that a random permutation of  $n$  symbols has no fixed points. HINT: Principle of inclusion and exclusion. (Your answer may involve a finite sum, which you don't need to simplify.)

*Solution.* Use inclusion-exclusion: let  $A_i$  be the set of permutations that  $\pi(i) = i$ .

$$\begin{aligned} p_n &= 1 - \mathbb{P}(\cup_{i=1}^n A_i) \\ &= 1 - \left( \sum_i \mathbb{P}(A_i) - \sum_{i,j} \mathbb{P}(A_i \cap A_j) + \cdots + (-1)^{n+1} \mathbb{P}(\cap_{i=1}^n A_i) \right) \\ &= 1 - n \cdot \frac{(n-1)!}{n!} + \binom{n}{2} \cdot \frac{(n-2)!}{n!} + \cdots + (-1)^n \binom{n}{n} \cdot \frac{1}{n!} \\ &= \sum_{p=0}^n (-1)^p \frac{1}{p!} \end{aligned}$$

(b)

Let  $S$  be a subset of  $\{1, 2, \dots, n\}$  of size  $k$ . Find the probability that the set of fixed points of a random permutation on  $n$  symbols is equal to  $S$ , and find the probability that a permutation has exactly  $k$  fixed points. HINT: If you didn't find the values  $p_j$  in part(a), you can still give answers for (b) expressed in terms of one or more  $p_j$ 's.

*Solution.*

$$\begin{aligned} \mathbb{P}(\{\text{fixed points}\} = S) &= \mathbb{P}(\pi(i) = i, \forall i \in S) \cdot \mathbb{P}(\pi(j) \neq j, \forall j \in S^c | \pi(i) = i, \forall i \in S) \\ &= \frac{(n-k)!}{n!} p_{n-k} \\ \mathbb{P}(k \text{ fixed points}) &= \binom{n}{k} \cdot \mathbb{P}(\{\text{fixed points}\} = S) \\ &= \frac{p_{n-k}}{k!} \end{aligned}$$

We get the second probability knowing that there are  $\binom{n}{k}$  many sets with  $k$  fixed points.

**(c)**

Show that as  $n$  tends to infinity, the distribution of the number of fixed points converges to a Poisson(1) distribution.

*Proof.*

$$\begin{aligned} \lim_{n \rightarrow \infty} \mathbb{P}(k \text{ fixed points}) &= \frac{1}{k!} \lim_{n \rightarrow \infty} \sum_{p=0}^{n-k} \frac{(-1)^p}{p!} \\ &= \frac{1}{k!} \sum_{p=0}^{\infty} \frac{(-1)^p}{p!} \\ &= \frac{e^{-1}}{k!} \\ &\sim \text{Poisson}(1) \end{aligned}$$

□

## Problem 2

Let  $\{S_n, n \geq 0\}$  be symmetric simple random walk, that is,  $S_n = \sum_{i=1}^n \xi_i$  with  $\xi_1, \xi_2, \dots$  i.i.d. satisfying  $\mathbb{P}(\xi_1 = 1) = \mathbb{P}(\xi_1 = -1) = \frac{1}{2}$ . Let  $T = \min\{n : S_n = 0\}$ , and write  $\mathbb{P}_a$  for probabilities when the walk starts at  $S_0 = a$ . By *basic probabilities* for  $\{S_n\}$  we mean probabilities of the form  $\mathbb{P}_0(S_n = k)$ ,  $\mathbb{P}_0(S_n \geq k)$ , or  $\mathbb{P}_0(S_n \leq k)$ , all of which corresponding to starting at  $S_0 = 0$ .

**(a)**

For  $a \geq 1, i \geq 1, n \geq 1$ , express  $\mathbb{P}_a(S_n = i, T \leq n)$  and  $\mathbb{P}_a(S_n = i, T > n)$  in terms of finitely many basic probabilities. HINT: Reflection principle.

*Solution.* Use reflection (reflect the part of path after the first approach at 0, with respect to 0), we have:

$$\mathbb{P}_a(S_n = i, T \leq n) = \mathbb{P}_a(S_n = -i) = \mathbb{P}_0(S_n = i + a)$$

Use conditional probability, we have:

$$\begin{aligned}
 \mathbb{P}_a(S_n = i, T > n) &= \mathbb{P}_a(S_n = i) \cdot \mathbb{P}_a(T > n | S_n = i) \\
 &= \mathbb{P}_a(S_n = i) \cdot (1 - \mathbb{P}_a(T \leq n | S_n = i)) \\
 &= \mathbb{P}_a(S_n = i) \cdot \left(1 - \frac{\mathbb{P}_a(T \leq n, S_n = i)}{\mathbb{P}_a(S_n = i)}\right) \\
 &= \mathbb{P}_a(S_n = i) - \mathbb{P}_a(S_n = i, T \leq n) \\
 &= \mathbb{P}_0(S_n = i - a) - \mathbb{P}_0(S_n = i + a)
 \end{aligned}$$

(b)

For  $a \geq 1$ ,  $i \geq 1$ ,  $n \geq 1$ , show that

$$\mathbb{P}_a(T > n) = \sum_{j=1-a}^a \mathbb{P}_0(S_n = j).$$

HINT: use (a) and look for cancellation

*Proof.*

$$\begin{aligned}
 \mathbb{P}_a(T > n) &= \sum_{i=1}^{a+n} \mathbb{P}_a(S_n = i, T > n) \\
 &= \sum_{i=1}^{a+n} \mathbb{P}_0(S_n = i - a) - \mathbb{P}_0(S_n = i + a) \\
 &= \sum_{i=1-a}^n \mathbb{P}_0(S_n = i) - \sum_{j=1+a}^n \mathbb{P}_0(S_n = j) \\
 &= \sum_{j=1-a}^a \mathbb{P}_0(S_n = j)
 \end{aligned}$$

□

(c)

You may take as given that  $\mathbb{P}_0(S_{2m} = 2j) \sim 1/\sqrt{\pi m}$  as  $m \rightarrow \infty$  for each fixed  $j \in \mathbb{Z}$ ; here  $\sim$  means that ratio converges to 1. Use this to find  $c$ ,  $\alpha$  such that  $\mathbb{P}_a(T > n) \sim c/n^\alpha$  as  $n \rightarrow \infty$ , where  $a > 0$ . Does  $c$  or  $\alpha$  depend on  $a$ ? HINT: It's enough to consider even  $n$  - why?

*Proof.* Assume  $n$  is even where  $n = 2m$ . For very large  $n$ , we have:

$$\begin{aligned}\mathbb{P}_a(T > 2m) &= \sum_{j=1-a}^a \mathbb{P}_0(S_{2m} = j) \\ &= \sum_{j \in A} \mathbb{P}_0(S_{2m} = j), \quad A = \{\text{even numbers in } \{1-a, 2-a, \dots, a\}\} \\ &\sim a \cdot \frac{1}{\sqrt{\pi m}} \\ &= \frac{a\sqrt{\frac{2}{\pi}}}{n^{1/2}}\end{aligned}$$

So we get  $c = a\sqrt{\frac{2}{\pi}}$  and  $\alpha = \frac{1}{2}$ , where  $c$  depends on  $a$ ,  $\alpha$  does not.

Now we assume  $n$  is odd, and we will prove the convergence by squeezing. First by inclusion, we have the inequality:

$$\mathbb{P}_a(T > n-1) \geq \mathbb{P}_a(T > n) \geq \mathbb{P}_a(T > n+1)$$

divide the expected limit:

$$\frac{\mathbb{P}_a(T > n-1)}{c/n^\alpha} \geq \frac{\mathbb{P}_a(T > n)}{c/n^\alpha} \geq \frac{\mathbb{P}_a(T > n+1)}{c/n^\alpha}$$

normalize both sides:

$$\frac{\mathbb{P}_a(T > n-1)}{c/(n-1)^\alpha} \cdot \left(\frac{n}{n-1}\right)^\alpha \geq \frac{\mathbb{P}_a(T > n)}{c/n^\alpha} \geq \frac{\mathbb{P}_a(T > n+1)}{c/(n+1)^\alpha} \cdot \left(\frac{n}{n+1}\right)^\alpha$$

Now, notice  $n-1$  and  $n+1$  are even, so if we let  $n$  go to infinity, both upper and lower bound above will converge to 1.  $\square$

### Problem 3

Let  $X, Y$  be independent standard normal  $(0, 1)$  random variables.

(a)

Find  $a$  for which  $U = X + 2Y, V = aX + Y$  are independent.

*Solution.* Note that  $U = (1, 2) \cdot (X, Y)^T, V = (a, 1) \cdot (X, Y)^T$ , and  $(X, Y)^T \sim \mathcal{N}(0, I)$ .  $(U, V)$  are normal vector, so  $U, V$  are independent if and only if  $\text{Cov}(U, V) = 0$ .

$$\begin{aligned}\text{Cov}(U, V) &= (1, 2) \cdot I \cdot (a, 1)^T \\ &= a + 2 \\ a &= -2\end{aligned}$$

**(b)**

Find  $\mathbb{E}(XY|X + 2Y = a)$  for all  $a \in \mathbb{R}$ . HINT: Use(a).

*Solution.* Note that  $X = \frac{U-2V}{5}$  and  $Y = \frac{2U+V}{5}$ . So the expectation turns into:

$$\frac{1}{25}\mathbb{E}(2U^2 - 3UV - 2V^2|U = a) = \frac{1}{25}(2a^2) - 3a \cdot \mathbb{E}(V) - 2 \cdot \mathbb{E}(V^2) = \frac{2a^2 - 10}{25}$$