## Fall 2014 Math 541b Exam

1. (a) Let $q_{x, y}$ be a Markov transition function, and $\pi_{x}$ a probability distribution on a finite state space $S$. Show that the Markov chain that accepts moves made according to $q_{x, y}$ with probability

$$
p_{x, y}=\min \left\{\frac{\pi_{y} q_{y, x}}{\pi_{x} q_{x, y}}, 1\right\}
$$

and otherwise remains at $x$, has stationary distribution $\pi_{x}$. Show that if $q_{x, y}$ and $\pi_{x}$ are positive for all $x, y \in S$ then the chain so described has unique stationary distribution $\pi_{x}$.
(b) Let $f(y)$ and $g(y)$ be two probability mass functions, both positive on $\mathbb{R}$. With $X_{1}$ generated according to $g$, consider the Markov chain $X_{1}, X_{2}, \ldots$ that for at stage $n \geq 1$ generates an independent observation $Y_{n}$ from density $g$, and accepts this value as the new state $X_{n+1}$ with probability

$$
\min \left\{\frac{f\left(Y_{n}\right) g\left(X_{n}\right)}{f\left(X_{n}\right) g\left(Y_{n}\right)}, 1\right\}
$$

and otherwise sets $X_{n+1}$ to be $X_{n}$. Prove that the chain converges in distribution to a random variable with distribution $f$.
(c) The accept/reject method. Let $f$ and $g$ be density functions on $\mathbb{R}$ such that the support of $f$ is a subset of the support of $g$, and suppose that there exists a constant $M$ such that $f(x) \leq M g(x)$. Consider the procedure that generates a random variable with distribution $g$, an independent random variable with the uniform distribution $U$ on $[0,1]$ and sets $Y=X$ when $U \leq f(X) / M g(X)$. Show that $Y$ has density $f$.
2. Let $f$ be a real valued function on $\mathbb{R}^{n}$, and $Z=f\left(X_{1}, \ldots, X_{n}\right)$ for $X_{1}, \ldots, X_{n}$ independent random variables.
(a) With $E^{(i)}(\cdot)=E\left(\cdot \mid X_{1}, \ldots, X_{i-1}, X_{i+1}, \ldots, X_{n}\right)$ show the following version of the Efron-Stein inequality

$$
\begin{equation*}
\operatorname{Var}(Z) \leq E\left(\sum_{i=1}^{n}\left(Z-E^{(i)} Z\right)^{2}\right) \tag{1}
\end{equation*}
$$

Hint: With $E_{i}(\cdot)=E\left(\cdot \mid X_{1}, \ldots, X_{i}\right)$, show that

$$
Z-E Z=\sum_{i=1}^{n} \Delta_{i} \quad \text { where } \quad \Delta_{i}=E_{i} Z-E_{i-1} Z
$$

compute the variance of $Z$ in this form, use properties of conditional expectation such as $E_{i}\left(E^{(i)}(\cdot)\right)=E_{i-1}(\cdot)$, and (conditional) Jensens' inequality.
(b) Letting $\left(X_{1}^{\prime}, \ldots, X_{n}^{\prime}\right)$ be an independent copy of $\left(X_{1}, \ldots, X_{n}\right)$, with

$$
Z_{i}^{\prime}=f\left(X_{1}, \ldots, X_{i-1}, X_{i}^{\prime}, X_{i+1}, \ldots, X_{n}\right)
$$

show that

$$
\operatorname{Var}(Z) \leq \frac{1}{2} E\left(\sum_{i=1}^{n}\left(Z-Z_{i}^{\prime}\right)^{2}\right)
$$

Hint: Express the right hand side of (1) in terms of conditional variances, and justify and use the conditional version of the fact that if $X$ and $Y$ are independent and have the same distribution then the variance of $X$ can be expresses in terms of $E(X-Y)^{2}$.

