Dynamic Valuation and Hedging of a Book

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- Unhedged Valuation

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- Analysis of Results

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- They define the book to be valued and hedged.

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• For stock positions in dollar terms of a(S(u), u) at time u and discount rate r the present value of such flows to date are

$$\sum_{t_i \leq t} c_i \left(S(t_i), t_i \right) e^{-rt_i} + \int_0^t \int_{-\infty}^\infty e^{-ru} a(S(u_{-}), u) \left(e^x - 1 \right) \mu(dx, du)$$

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 where µ(dx, du) is the integer valued random measure associated with the finite variation jumps in the log price relative of the stock price. • For valuation we follow the \mathcal{G} – expectations approach introduced in Peng (2006) with valuations defined by nonlinear expectations that are unique viscosity solutions to equations of the form

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• The result is a nonlinear valuation at time t of the future cash flows yet to be realized, when the spot at time t is at level S.

Spatially Inhomogeneous Compensator

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We write

$$\nu(S, x) dx dt = \begin{pmatrix} \frac{c_p(S)}{x} \exp\left(-\frac{x}{b_p(S)}\right) \mathbf{1}_{x>0} \\ + \frac{c_n(S)}{|x|} \exp\left(-\frac{|x|}{b_n(S)}\right) \mathbf{1}_{x<0} \end{pmatrix} dx dt$$

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• The spatially inhomogeneous compensating measure is then

$$\nu(S,A) = \int_{A} \nu(S,x) \, dx$$

• The law of one price coupled with no arbitrage implies that the value function V(X) satisfies

$$V(aX + bY) = aV(X) + bV(Y).$$

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- Hence the valuation function is linear.
- The set of acceptable risks are those with a positive value and this is a very large convex set.
- Useful as it may be, it renders optimization useless and defines risk acceptability too generously.

Graph of Arbitrages, Positive Alpha, and Acceptable Opportunities



• When risk acceptability is contracted from a half space to a proper convex cone containing the nonnegative outcomes that are certainly acceptable at zero cost, then the acceptable risks become all outcomes X satisfying

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Conservative valuation in two price economies

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- are acceptable.
- The best and bid and ask prices are

$$b(X) = \inf_{Q \in \mathcal{M}} E^{Q}[X]$$

$$a(X) = \sup_{Q \in \mathcal{M}} E^{Q}[X]$$

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- The maximization objective for risk exposure evaluation should be concave.
- As a consequence the set of outcomes with a positive value are a convex set but not in general a convex cone.
- We now ask why one should scale value with the scale of the outcome?
- If the value is market based and contemplated positions are small relative to the size of the market then the value of twice the outcome should be twice the value.

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- Let $\Psi(u)$ be a concave distribution function on the unit interval and evaluate the value of outcome X with distribution function F(x) and density f(x) as

$$b(X) = \int x \Psi'(F(x)) f(x) dx$$

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• $\Psi'(u)$ for u near zero lifts the weighting on losses while $\Psi'(u)$ for u near unity reduces the weighting on gains.

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• The lower price formed from distorted expectations is associated with \mathcal{M} the set of supporting measures being all measures Q such that

 $Q(A) \leq \Psi(P(A))$, for all A.

Bid Ask Distortions



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- All outcomes are bounded away from zero on sets of finite measure.
- On sets of infinite measure outcomes converge towards zero reflecting the view that only nothing happens all the time.

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• The upper value is given by

$$\mathsf{a}(X) = -\int_0^\infty G^-\left(\mu\left(X < - \mathsf{a}
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ight)\mathsf{d}\mathsf{a} + \int_0^\infty G^+\left(\mu\left(X > \mathsf{a}
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• The lower valuation using measure distortions is

$$b(X) = \inf_{\widetilde{\nu} \in \mathcal{M}} \int X(\omega) \widetilde{\nu}(d\omega)$$
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• where $\widetilde{\nu} \in \mathcal{M}$ just if

 ${{\mathcal G}^-}(\nu(A)) \leq \widetilde{\nu}(A) \leq {{\mathcal G}^+}\left(\nu(A)\right)\text{, for all }A \text{ with }\nu(A) < \infty.$

Measure Distortions



Hedged Valuation

• For two measure increasing distortions G^+ , G^- concave/convex, above/below the identity and hedge policy a(S, t) we define the operator

$$\begin{aligned} \mathcal{G}(V) &= -\int_0^\infty G^+ \left(\nu(S, \left[\begin{array}{c} a(S,t) \left(e^x - 1 \right) \\ + V(Se^x,t) - V(S,t) \end{array} \right]^- > w \right) dw \\ &+ \int_0^\infty G^- \left(\nu(S, \left[\begin{array}{c} a(S,t) \left(e^x - 1 \right) \\ + V(Se^x,t) - V(S,t) \end{array} \right]^+ > w \right) dw \end{aligned}$$

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• The hedge selection defines

$$a(S,t) = \arg \max_{a} \left(-\int_{0}^{\infty} G^{+} \left(\nu(S, \left[\begin{array}{c} a(e^{x}-1) \\ +V(Se^{x},t) - V(S,t) \end{array} \right]^{-} > w \right) dw \\ +\int_{0}^{\infty} G^{-} \left(\nu(S, \left[\begin{array}{c} a(e^{x}-1) \\ +V(Se^{x},t) - V(S,t) \end{array} \right]^{+} > w \right) dw \\ +V(Se^{x},t) - V(S,t) = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}{c} e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \\ e^{x} - 1 \end{array} \right]^{+} = \left[\begin{array}[c] e^{x} - 1 \end{array} \right]^{+} = \left[$$

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with

$$\begin{aligned} X(t) &= \int_{(0,t]\times\mathbb{R}^k\setminus\{0\}} ((\exp\circ I)(x) - \mathbf{1}_k)\nu(X(s), s, x)dxds \\ &+ \int_{(0,t]\times\mathbb{R}^k\setminus\{0\}} ((\exp\circ I)(x) - \mathbf{1}_k)\widetilde{N}(dx \times ds) \end{aligned}$$

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$$X(t) = \int_{(0,t]\times\mathbb{R}^k\setminus\{0\}} ((\exp\circ I)(x) - \mathbf{1}_k)\nu(X(s), s, x)dxds + \int_{(0,t]\times\mathbb{R}^k\setminus\{0\}} ((\exp\circ I)(x) - \mathbf{1}_k)\widetilde{N}(dx \times ds)$$

• where $(\exp \circ I)(x) = (e^{x_1}, \cdots, e^{x_k})^T$, $\mathbf{1}_k$ is the k dimensional vector with all entries unity, and $\widetilde{N}(dx \times ds)$ is a compensated jump

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• On this probability space let χ be a terminal random variable and let B_t be a lower prudential valuation for χ associated with a driver function g(t, z) defined on $(0, T] \times \mathcal{L}^2(\nu(X(t), t, x)dx)$ that is positive, homogeneous and convex in z.

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- The Backward Stochastic Partial Integro-Differential Equation (BSPIDE) solves for (B_t, Z_t) for t < T the equation

$$B_t + \int_t^T g(s, Z_s) ds + \int_{(0,t] imes \mathbb{R}^k \setminus \{0\}} Z_s(x) \widetilde{N}(ds imes dx) = \chi.$$

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 \bullet and V solves the semilinear partial integro-differential equation

$$\begin{split} \dot{V} + \mathcal{K}V(t, x) - g(t, \mathcal{D}V_{t,x}) \\ &= 0 \\ \mathcal{K}V(t, x) = d_t^T \nabla V \\ &+ \int_{\mathbb{R}^k \setminus \{0\}} (\mathcal{D}V_{t,x} - \nabla V(t, x)^T x (e^y - 1)) \nu(dy) \\ \mathcal{D}V_{t,x} = V(t, xe^y) - V(t, x) \\ d_t &= \int_{\mathbb{R}^k \setminus \{0\}} x (e^y - 1) \nu(x, t, dy) \end{split}$$

Valuations and Risk Acceptability

• The lower prudential value is related to risk acceptability via

$$B_t = \inf_{Q \in \mathcal{S}^g} E^Q \left[\chi | \mathcal{F}_t \right].$$
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• Here S^g consists of all measures Q for which the representation M of the stochastic logarithm satisfies a condition related to g.

$$\begin{array}{lll} \xi &=& \displaystyle \frac{dQ}{dP} \\ \xi &=& \displaystyle \mathcal{E}(M), \ \mathcal{E} \ \text{is stochastic exponential} \\ M &=& \displaystyle \int_{(0,T]\times\mathbb{R}^k\setminus\{0\}} H_s(y)\widetilde{N}(ds\times dy) \end{array}$$

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• with

$$\int z(y)H_s(y)\nu(X(s),s,y)dy \leq g(s,z), \ z \in \mathcal{L}^2\left(\nu(X(s),s,y)dy\right).$$

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$$W + GW = 0$$

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The Valuation Driver

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- The distorted variation we work with is

$$\begin{aligned} \mathcal{G}W &= -\int_0^\infty G^+ \left(\nu(X, [W(Xe^x, t) - W(X, t)]^- > w \right) dw \\ &+ \int_0^\infty G^- \left(\nu\left(X, [W(Xe^x, t) - W(X, t)]^+ > w \right) \right) dw. \end{aligned}$$

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• The variation on the other hand is

$$\mathcal{V}(W) = -\int_0^\infty \left(\nu(X, [W(Xe^x, t) - W(X, t)]^- > w \right) dw$$

+
$$\int_0^\infty - \nu \left(X, [W(Xe^x, t) - W(X, t)]^+ > w \right) dw$$

The Valuation Driver II

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• and observe that

$$\mathcal{G}W = \mathcal{V}W - \int_0^\infty \Gamma^+ \left(\nu(X, [W(Xe^x, t) - W(X, t)]^- > w \right) dw$$

$$- \int_0^\infty \Gamma^- \left(\nu\left(X, [W(Xe^x, t) - W(X, t)]^+ > w \right) \right) dw.$$

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• The semilinear equation in W is on noting $\mathcal{K}W = \mathcal{V}W$ that

$$W + \mathcal{K}W - g(t, \mathcal{D}W_{t,x}) = 0$$

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• We now write that

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• Furthermore we have that

$$e^{-rt}V(t,X(t)) = \inf_{\mathbb{Q}\in\mathcal{S}^g} E^{\mathbb{Q}}[e^{-rT}\chi|\mathcal{F}_t].$$

Illustrate on a Vanilla Book of SPY options

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- But for a genuine Markov context we must establish a dependence of the law of motion on observables.
- It is unreasonable to suppose that *SPY* return distributions over long periods used in estimation will show a dependence of such on the level of the index.
- We consider instead a dependence on the ratio of the index level to a geometrically weighted average of past prices.

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- We then bucket the returns based on the level of the ratio.
- The bilateral gamma model is then estimated separately in each bucket using digital moment estimation.
- From the estimated parameters one may evaluate the return drift separately in each bucket.

Drift and the Ratio

• The Figure presents the dependence observed for the drift as a function of the ratio.



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- At all points including the coarse grid the actual parameters are taken from the coarse grid candidate values as a smooth function using a Gaussian Kernel smoother.
- This avoids having to describe up front a functional form for the dependence.
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- For arbitrary Z levels, we use the closest Z value on the fine grid where the density has been precomputed.
- This minimizes the number of Fourier inversions to be computed.
- The logarithm of the thousand densities evaluated at the observed returns are summed to construct the log likelihood to be maximized.

Parameter Dependence on Spot Average Ratio

• The Figure presents the dependence of the four bilateral gamma parameters on the spot average ratio.

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Dependence of Expected Return on Spot Average Ratio

• The Figure show the dependence of the expected return in basis points on the level of the spot average ratio.

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- The Figure show the dependence of the expected return in basis points on the level of the spot average ratio.
- The drift though positive falls with the level of the ratio.



Markov Model for the Average, Y, and the Spot Average Ratio, Z

• In a continuous formulation the average Y(t) may be constructed from the Spot prices S(t) by

$$Y(t) = \theta \int_{-\infty}^{t} e^{-\theta(t-u)} S(u) du.$$

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• The dynamics for the ratio are given by

 $dZ = -\theta Z(t) \left(Z(t) - 1 \right) dt Z(t) \left(\left(e^{x} - 1 \right) * \mu(dx, du) \right)$

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• The compensator for Z is given by

$$\nu(Z, dx) dt = \begin{pmatrix} \frac{c_p(Z)}{x} \exp\left(-\frac{x}{b_p(Z)}\right) \mathbf{1}_{x>0} \\ + \frac{c_n(Z)}{|x|} \exp\left(-\frac{|x|}{b_n(Z)}\right) \end{pmatrix} dxdt$$

• Hence Z(t), Y(t) is a two dimensional Markov process and we seek the value and hedge policy functions

V(Z(t), Y(t), t)a(Z(t), Y(t), t)

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to be solved for on a three dimensional grid for Z, Y and time t.
We do this on an 18 core Alienware Machine.

• We solve

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$$\begin{aligned} \mathcal{G}(V) \\ &= -\int_0^\infty G^+ \left(\nu(Z, \left[\begin{array}{c} a(Z, Y, t) (e^x - 1) \\ + V(Ze^x, Y, t) - V(Z, Y, t) \end{array} \right]^- > w) \right) dw \\ &+ \int_0^\infty G^- \left(\nu(Z, \left[\begin{array}{c} a(Z, Y, t) (e^x - 1) \\ + V(Ze^x, Y, t) - V(Z, Y, t) \end{array} \right]^+ > w) \right) dw. \end{aligned}$$

 $V(Z, Y, 0) = C_T(YZ)$ $V(Z, Y, t_i) = V(Z, Y, t_{i-}) + c_i(Y(t_{i-})Z), t_i > 0.$

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The policies locally maximize G(V) over choices for hedge positions a.

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- We also take a 40 point non-uniform grid on the ratio between 0.9 and 1.1.
- The book of claims hedged have maturities under a year.
- The backward recursion is implemented on 20 time steps.

 The integral of the compensator v(Z, x)dx over sets A is accomplished by simulation.

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- Values and Hedge positions are smoothed at each time step using Gaussian Process Regression.

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- Summing the spot dependent option payoffs times positions delivers the functions $c_i(S)$.
- There were 29 maturities below a year in the option book ranging from one to 358 days.
Sample Claims By Maturity

• The Figure presents graphs of state contingent claims in millions of dollars for a sample of the 29 maturities



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• We present a sample of graphs showing how hedge positions vary with Z for different levels of Y and time.

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Value Function By Z and Y

• We present four graphs showing the nonlinear value as a function of Z for three settings on Y for now, three, six, and nine months in.



Hedge Position Regression Coefficients

• The hedge positions were regressed on the delta, gamma, and levels of Z, Y, Z^2 , Y^2 separately for each level of time maturity.

Hedge Position Regression Coefficients

- The hedge positions were regressed on the delta, gamma, and levels of Z, Y, Z², Y² separately for each level of time maturity.
- The Table shows results for a sample of times to maturity.

			Regressing	g Hedge Po			
Time Left	Delta	Gamma	Z	Y	Z^2	Y^2	RSQ
1	1.1293	29.9751	1723.6023	195.7961	-8.3502	-0.3197	0.9302
0.9	1.0964	36.8843	318.8334	48.4325	-1.5291	-0.0785	0.9243
0.8	1.4404	40.2417	95.7524	12.5620	-0.4641	-0.0205	0.9200
0.7	2.1472	66.0671	19.6692	3.4710	-0.0950	-0.0057	0.9080
0.6	2.0326	62.7417	8.7496	1.6764	-0.0425	-0.0027	0.9007
0.5	2.3538	73.7335	4.8765	0.8572	-0.0237	-0.0014	0.8791
0.4	5.6210	191.1055	2.6347	0.4515	-0.0129	-0.0007	0.8611
0.3	6.9205	197.6278	3.2125	0.4428	-0.0158	-0.0007	0.8342

• The Table shows the corresponding t-stats.

			T-Stats			
Time Left	Delta	Gamma	Z	Y	Z^2	Y^2
1	17.77	10.68	16.77	73.64	-16.24	-71.09
0.9	15.46	11.71	11.67	68.52	-11.20	-65.67
0.8	18.08	11.29	13.40	67.87	-12.99	-65.32
0.7	22.23	15.23	9.77	66.50	-9.43	-64.38
0.6	19.86	13.64	8.75	64.65	-8.50	-62.60
0.5	19.28	13.47	8.65	58.71	-8.42	-56.99
0.4	18.96	14.29	8.36	55.32	-8.17	-53.76
0.3	21.73	13.70	9.45	50.24	-9.32	-49.09

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Conclusion

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