

# A Berry-Esseen Theorem for the Lightbulb Process

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## The Lightbulb Process

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In the lightbulb process of Rao, Rao and Zhang, on days  $r = 1, \dots, n$ , out of  $n$  lightbulbs, all initially off, exactly  $r$  bulbs, selected uniformly and independent of the past, have their status changed from off to on, or vice versa.

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Their histograms indicated that  $X$  is asymptotically normal, but left the question open. 'Out of the box' central limit theorems seem to not apply easily to this process.

## Mean and Variance

For  $s = 1, \dots, n$ , let

$$\lambda_{n,1,s} = 1 - \frac{2s}{n} \quad \text{and} \quad \lambda_{n,2,s} = 1 - \frac{4s}{n} + \frac{4s(s-1)}{n(n-1)},$$

and  $\lambda_{n,b,\mathbf{s}} = \prod_{r=1}^n \lambda_{n,b,s_r}$ . Then with switch pattern  $\mathbf{s} = (s_1, \dots, s_n)$ ,

$$EX_{\mathbf{s}} = \frac{n}{2} (1 - \lambda_{n,1,\mathbf{s}})$$

and

$$\text{Var}(X_{\mathbf{s}}) = \frac{n}{4} (1 - \lambda_{n,2,\mathbf{s}}) + \frac{n^2}{4} (\lambda_{n,2,\mathbf{s}} - \lambda_{n,1,\mathbf{s}}^2).$$

## Berry Esseen Theorem for Lightbulb

**Theorem 1** *Let  $X$  be the number of bulbs on at the terminal time  $n$ , an even integer, and  $\mu = n/2$  and  $\sigma^2 = \text{Var}(X)$ . Then*

$$\sup_{z \in \mathbb{R}} \left| P \left( \frac{X - \mu}{\sigma} \leq z \right) - P(Z \leq z) \right| \leq \frac{n}{2\sigma^2} \Psi_0 + 1.64 \frac{n}{\sigma^3} + \frac{2}{\sigma}$$

where  $Z$  is standard normal and

$$\Psi_0 \leq \frac{1}{2\sqrt{n}} + \frac{1}{2n} + e^{-n/2} \quad \text{for } n \geq 6.$$

Yields a bound of order  $O(n^{-1/2})$  as  $n \rightarrow \infty$ .



## Composition Markov chains of multinomial type

Considered by Zhou and Lange, such chains in general are based on a  $d \times d$  Markov transition matrix  $P$  which describes the transition of a single particle in a system of  $n$  identical particles, a subset of which is selected uniformly to undergo transition at each time step according to  $P$ .

Very complete spectral decompositions of such chains may be available; this decomposition becomes essential to the calculation of the Berry Esseen bound for this problem.

Ehrenfest chains, Hoare-Rahman chains, Kimura's model for DNA base-pair substitution.

## Size Bias Coupling

Let  $X$  be nonnegative with finite mean  $\mu$ . Recall  $X^s$  has the  $X$ -sized biased distribution if

$$E[Xf(X)] = \mu E[f(X^s)]$$

for all smooth  $f$ .

If  $X = \sum_{i=1}^n X_i$  is the sum of exchangeable indicators, then one may form  $X^s$  by picking  $i$  uniformly and summing the variables having distribution

$$\mathcal{L}(X_1^i, \dots, X_n^i) = \mathcal{L}(X_1, \dots, X_n | X_i = 1)$$

## Size biased Coupling

With  $W = (X - \mu)/\sigma$ , and  $W^s = (X^s - \mu)/\sigma$ ,

$$\begin{aligned} & E(h(W) - Eh(Z)) \\ &= E(f'(W) - Wf(W)) \\ &= E\left(f'(W) - \frac{\mu}{\sigma}(f(W^s) - f(W))\right) \\ &= E\left(f'(W)\left(1 - \frac{\mu}{\sigma}(W^s - W)\right) - \frac{\mu}{\sigma} \int_0^{W^s - W} (f'(W + t) - f'(W))dt\right). \end{aligned}$$

When monotone  $W^s \geq W$ .

## Concentration Inequality

Size bias version of exchangeable pair concentration inequality of Shao and Su (2005).

**Lemma 1** *Let  $X$  be a nonnegative random variable with mean  $\mu$  and variance  $\sigma^2$ , both finite and positive, and let  $X^s$  be given on the same space as  $X$ , having the  $X$  size biased distribution and satisfying  $X^s \geq X$ . Then with  $W = (X - \mu)/\sigma$  and  $W^s = (X^s - \mu)/\sigma$ , for any  $z \in \mathbb{R}$  and  $a \geq 0$ ,*

$$\frac{\mu}{\sigma} E(W^s - W) \mathbf{1}_{\{W^s - W \leq a\}} \mathbf{1}_{\{z \leq W \leq z+a\}} \leq a.$$

## Monotone Size Biased Coupled Pair

**Theorem 2** *Let  $X$  be a nonnegative random variable with mean  $\mu$  and variance  $\sigma^2$ , both finite and positive, and let  $X^s$  be given on the same space as  $X$ , with the  $X$  size biased distribution, satisfying  $X \leq X^s \leq X + B$  for some  $B > 0$ . Then with  $W = (X - \mu)/\sigma$ , we have*

$$\sup_{z \in \mathbb{R}} |P(W \leq z) - P(Z \leq z)| \leq \frac{\mu}{\sigma^2} \Psi + 0.82 \frac{\delta^2 \mu}{\sigma} + \delta,$$

where

$$\Psi = \sqrt{\text{Var}(E(X^s - X|X))} \quad \text{and} \quad \delta = B/\sigma.$$

## Size biased coupling: Lightbulb

The number of bulbs on at time  $n$  is

$$X = \sum_{k=1}^n X_k,$$

where  $X_k$  is the indicator that bulb  $k$  is on. Need to select bulb at random, and turn it on if not already, keeping correct conditional distribution, and without greatly affecting the status of the other bulbs.

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With  $X_{rk}$  the switch variable of bulb  $k$  at time  $r$ , we have

$$X_k = \left( \sum_{r=0}^n X_{rk} \right) \bmod 2.$$

## Size biased coupling: Lightbulb

Let  $n = 2m$  and  $\mathbf{X} = \{X_{rk}, r, k = 1, \dots, n\}$ . For given bulb  $i$ , if  $X_i = 1$  then set  $\mathbf{X}^i = \mathbf{X}$ .

Otherwise  $J^i \sim \mathcal{U}\{j : X_{n/2,j} = 1 - X_{n/2,i}\}$ , independent of  $\{X_{rk} : r \neq n/2, k = 1, \dots, n\}$ . Let  $\mathbf{X}^i$  have components

$$X_{rk}^i = \begin{cases} X_{rk} & r \neq n/2 \\ X_{n/2,k} & r = n/2, k \notin \{i, J^i\} \\ X_{n/2,J^i} & r = n/2, k = i \\ X_{n/2,i} & r = n/2, k = J^i. \end{cases}$$



## Size biased coupling: Lightbulb

In words, if  $X_i = 0$  then uniformly select a  $j$  whose switch variable  $X_{n/2,j}$  at stage  $n/2$  is opposite to  $X_{n/2,i}$ , and interchange.

## Size biased coupling: Lightbulb

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If  $I$  is the uniformly chosen index, and  $J$  the index of the bulb with opposite status at stage  $n/2$ , then

$$X^s - X = 2\mathbf{1}_{\{X_I=0, X_J=0\}}.$$

The coupling is monotone and bounded.

## Conditional Variance Calculation, $n = 2m$

With  $\mathcal{F}$  the  $\sigma$ -algebra generated by the switch variables  $\mathbf{X}$ ,

$$E\left(X^s - X \mid \mathcal{F}\right) = \frac{4}{n^2} \sum_{i \neq j} \mathbf{1}_{\{X_i=0, X_j=0, X_{n/2,i}=0, X_{n/2,j}=1\}}.$$

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The variance calculation requires the computation of (four) joint probabilities such as  $g_{2,2,n,(1,\dots,n),n/2}(g_{\alpha,\beta,n,s,l})$

$$P(X_{i_1} = 0, X_{j_1} = 0, X_{i_2} = 0, X_{j_2} = 0, \\ X_{n/2,i_1} = 0, X_{n/2,j_1} = 1, X_{n/2,i_2} = 0, X_{n/2,j_2} = 1).$$

Conditioning on switches at stage  $n/2$  they later become 'initial conditions.'

## Spectral Decomposition

For a given subset of  $b$  of the  $n$  bulbs, the eigenvalues of one, and multiple steps of the chain are given by

$$\lambda_{n,b,s} = \sum_{t=0}^b \binom{b}{t} (-2)^t \frac{(s)_t}{(n)_t} \quad \text{and} \quad \lambda_{n,b,s} = \prod_{r=1}^k \lambda_{n,b,s_r},$$

where  $(n)_k = n(n-1)\cdots(n-k+1)$  denotes the falling factorial, and the empty product is 1.

## Spectrum of Composition Markov Chains of Multinomial Type

In the case of the lightbulb chain there are  $d = 2$  states and the transition matrix  $P$  of a single bulb is given by

$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}.$$

With  $b \in \{0, 1, \dots, n\}$  let  $P_{n,b,s}$  be the  $2^b \times 2^b$  transition matrix of a subset of size  $b$  of the  $n$  total lightbulbs when  $s$  of the  $n$  bulbs are selected uniformly to be switched.

Letting  $P_{n,0,s} = 1$  for all  $n$  and  $s$ , and  $I_2$  the  $2 \times 2$  identity matrix, for  $n \geq 1$  the matrix  $P_{n,b,s}$  is given recursively by

$$P_{n,b,s} = \frac{s}{n} (P \otimes P_{n-1,b-1,s-1}) + \left(1 - \frac{s}{n}\right) (I_2 \otimes P_{n-1,b-1,s}).$$

## Spectral Decomposition

Transition matrices are simultaneously diagonalizable by

$$P_{n,b,s} = \otimes^b T^{-1} \Gamma_{n,b,s} \otimes^b T,$$

where  $\Gamma_{n,b,s} = \text{diag}(\lambda_{n,a_1,s}, \dots, \lambda_{n,a_{2^b},s})$ , e.g.,

$\mathbf{a}_1 = (0, 1)$ ,  $\mathbf{a}_2 = (0, 1, 1, 2)$  and  $\mathbf{a}_3 = (0, 1, 1, 2, 1, 2, 2, 3)$ ,

and

$$T = \frac{1}{\sqrt{2}} \begin{bmatrix} 1 & 1 \\ -1 & 1 \end{bmatrix}.$$

# Spectral Decomposition

For calculation of  $g_{2,2,n,(1,\dots,n),n/2}$ ,

$$P(X_{i_1} = 0, X_{j_1} = 0, X_{i_2} = 0, X_{j_2} = 0, \\ X_{n/2,i_1} = 0, X_{n/2,j_1} = 1, X_{n/2,i_2} = 0, X_{n/2,j_2} = 1),$$

spectral decomposition yields

$$g_{2,2,n,\mathbf{s},l} = \frac{1}{16} (1 - 2\lambda_{n,2,\mathbf{s}_l} + \lambda_{n,4,\mathbf{s}_l}) \frac{(s_l)_2 (n - s_l)_2}{(n)_4},$$

where

$$\mathbf{s}_l = (s_1, \dots, s_{l-1}, s_{l+1}, \dots, s_n).$$



## Spectral Decomposition

Now turning to  $\lambda_{n,4,s}$ , for  $n \geq 4$  consider the fourth degree polynomial

$$f_4(x) = 1 - \frac{8x}{n} + \frac{24(x)_2}{(n)_2} - \frac{32(x)_3}{(n)_3} + \frac{16(x)_4}{(n)_4}, \quad 0 \leq x \leq n.$$

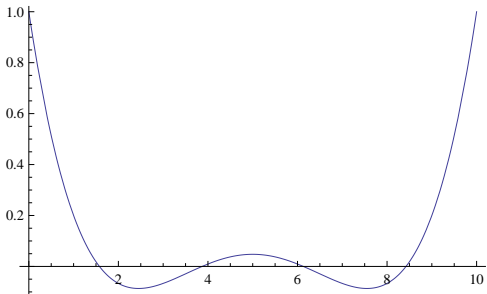
It can be checked that the four roots of  $f_4(x)$  are given by

$$x_{1\pm} = \frac{n \pm \sqrt{\sqrt{2}\sqrt{3n^2 - 9n + 8} + 3n - 4}}{2},$$

with  $0 < x_{1-} < x_{2-} < x_{2+} < x_{1+} < n$  where

$$x_{2\pm} = \frac{n \pm \sqrt{-\sqrt{2}\sqrt{3n^2 - 9n + 8} + 3n - 4}}{2},$$

$$f_4(x)$$



## Need Bounds to compute Conditional Variance

$$|f_4(x)| \leq \begin{cases} \frac{6}{(n-3)^2} & \text{for } x \in [x_{1-}, x_{1+}] \\ f_2^2(x) & \text{for } x \notin [x_{1-}, x_{1+}], x \in [0, n]. \end{cases}$$

where

$$f_2(x) = 1 - \frac{4x}{n} + \frac{4(x)_2}{(n)_2}, \quad 0 \leq x \leq n.$$

$$\lambda_{n,4,s}$$

$$|\lambda_{n,4,s}| = \prod_{s=0}^{\lfloor x_{1-} \rfloor} |\lambda_{n,4,s}| \prod_{s \in \mathbf{u}} |\lambda_{n,4,s}| \prod_{s=\lceil x_{1+} \rceil}^n |\lambda_{n,4,s}|$$

Obtain

$$|\lambda_{n,4,s}| \leq \frac{1}{2} e^{-n} \quad \text{for } n \geq 6,$$

for  $s = (1, 2, \dots, n/2 - 1, n/2 + 1, \dots, n)$ .

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Yields a bound of order  $O(n^{-1/2})$  as  $n \rightarrow \infty$ .

# Stein's Method for the Lightbulb Process

I hope that the talk was clear and illuminating



To Louis:

Thank you for all your hard work, and Happy Birthday!